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## Two Results on Jump Processes

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1. Introduction. Let  $(\Omega, \underline{F}, P)$  be a complete probability space, and  $X = (X_t)_{t \geq 0}$  a jump process, i.e. all its trajectories are r.c.l.l. (right-continuous and with left limits) step functions and have only finitely many jumps in every finite interval. Denote by  $(T_n)_{n \geq 1}$  the successive jump times of  $X$ , and by  $(\Delta_n)_{n \geq 1}$  the successive jump sizes of  $X$ . By convention we have  $T_0 = 0$  and  $\Delta_0 = X_0$ . Then  $X$  can be written as

$$X = X_0 + \sum_{n=1}^{\infty} \Delta_n I_{[T_n, \infty[}$$

and we have:

- 1)  $T_n \uparrow \infty$ ;
- 2) For all  $n \geq 0$ ,  $T_n < \infty \Rightarrow T_n < T_{n+1}$ ;
- 3) For all  $n \geq 1$ ,  $\Delta_n \neq 0 \Rightarrow T_n < \infty$ .

Denote by  $\underline{F} = (\underline{F}_t)_{t \geq 0}$  the natural filtration of  $X$ :

$$\underline{F}_t = \sigma\{X_s, s \leq t, \underline{N}\},$$

where  $\underline{N}$  is the family of  $P$ -null sets. It is well-known (see [3], [5] and [7]) that  $\underline{F}$  is right-continuous, so  $\underline{F}$  satisfies the usual conditions, and we have for any stopping time  $T$

$$\underline{F}_T = \sigma\{X^T, \underline{N}\}, \quad \underline{F}_{T-} = \sigma\{T, X^{T-}, \underline{N}\} \quad (1)$$

in particular, for all  $n \geq 1$

$$\underline{F}_{T_n} = \sigma\{\Delta_0, T_1, \Delta_1, \dots, T_n, \Delta_n, \underline{N}\}, \quad \underline{F}_{T_n-} = \sigma\{\Delta_0, T_1, \Delta_1, \dots, T_n, \underline{N}\} \quad (2)$$

Denote by  $\mu$  the jump measure induced by  $X$ :

$$\mu(dt, dx) = \sum_{n=1}^{\infty} \mathcal{E}_{(T_n, \Delta_n)}(dt, dx) I_{[T_n < \infty]}$$

where  $\mathcal{E}_a$  is the unite measure concentrating at point  $a$ , and by  $\nu$  the predictable dual projection of  $\mu$ . According to Jacod[7], we have

$$\nu(dt, dx) = \sum_{n=1}^{\infty} \frac{P(T_n \in dt, \Delta_n \in dx \mid \mathcal{F}_{T_{n-1}})}{P(T_n \geq t \mid \mathcal{F}_{T_{n-1}})} I_{\llbracket T_{n-1}, T_n \rrbracket} \quad (3)$$

The law of  $X$  is determined uniquely by that of  $(T_n, \Delta_n)_{n \geq 0}$  and by  $\nu$  as well. So it is natural to characterize the properties of  $X$  by the behaviour of  $(T_n, \Delta_n)_{n \geq 0}$  or  $\nu$ . In this note we show two simple but interesting results of this type.

We introduce another useful notations. Put

$$\Lambda_t = \nu([0, t] \times \mathbb{R}), \quad a_t = \Delta \Lambda_t = \nu(\{t\} \times \mathbb{R}).$$

It is easy to see that  $(\Lambda_t)$  is the predictable dual projection of the simple point process  $N = \sum_{n=1}^{\infty} I_{\llbracket T_n, \infty \rrbracket}$ ,  $(a_t)$  is the predictable projection of  $I_D$ , where  $D = [\Delta X \neq 0] = \bigcup_{n=1}^{\infty} \llbracket T_n \rrbracket$  is the set of the jumps of  $X$ , and  $J = [a \neq 0]$  is the predictable support of  $D$ . Suppose that on  $\{T_n < \infty\}$

$$P(\Delta_n \in dx \mid \mathcal{F}_{T_n^-}) = G_n(dx; \Delta_0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n) \text{ a.s.}$$

Then we have

$$\begin{aligned} \nu(dt, dx) &= G(t, dx) d\Lambda_t, \\ G(t, dx) &= \sum_{n=1}^{\infty} G_n(dx; \Delta_0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, t) I_{\llbracket T_{n-1}, T_n \rrbracket}(t) \end{aligned} \quad (4)$$

Our first result is concerned with the predictable representation property. We recall that  $X$  (or  $\mathcal{F}$ ) has the predictable representation property if there exists a  $\mathcal{F}$ -local martingale  $M$  such that every  $\mathcal{F}$ -local martingale  $L$ , with  $L_0 = 0$ , can be represented as a predictable stochastic integral  $H.M$ . In [4], under the assumption that  $\mathcal{F}$  is quasi-left-continuous we showed that  $X$  has the predictable representation property if and only if for every  $n \geq 1$ ,  $\Delta_n$  is a.s. a measurable function of  $(\Delta_0, T_1, \Delta_1, \dots, T_n)$ , or equivalently,  $\mathcal{F}$  is exactly the natural filtration of the simple point process  $\Delta_0 + N$ . But we know (see Chow and Meyer[1]) that the process  $\Delta_0 + N$  has always the predictable representation property. It is not reasonable to assume that the natural filtration  $\mathcal{F}$  is quasi-left-continuous. Now we get the general result as follows.

Theorem 1. The following statements are equivalent:

1°  $X$  has the predictable representation property;

2° For every  $n \geq 1$ , there exist two Borel functions  $f_n^{(i)}(x_0, t_1, x_1, \dots, t_{n-1}, x_{n-1}, t_n)$ ,  $i = 1, 2$ , such that on the set  $\{T_n < \infty\}$  we have

$$1) \Delta_n = f_n^{(1)}(\Delta_0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n) \text{ a.s.} \quad \text{on } \{a_{T_n} < 1\},$$

$$2) \Delta_n \in \{f_n^{(i)}(\Delta_0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n), i = 1, 2\} \text{ a.s. on } \{a_{T_n} = 1\}.$$

In other words, the conditional distribution of  $\Delta_n$  with respect to  $\mathbb{F}_{T_n-}$  on the set  $\{T_n < \infty\}$  is a two-valued discrete distribution, furthermore, it reduces to an

unite one on the set  $\{a_{T_n} < 1\}$ ;

3° There exist four predictable processes  $(c_t^{(i)}), (\alpha_t^{(i)}), i = 1, 2$ , with  $c_t^{(1)} \geq 0, c_t^{(2)} \geq 0, c_t^{(1)} + c_t^{(2)} = 1$ , such that

$$v(dt, dx) = G(t, dx) d\Lambda_t, \quad G(t, dx) = c_t^{(1)} \mathcal{E}_{(\alpha_t^{(1)})}(dx) + c_t^{(2)} \mathcal{E}_{(\alpha_t^{(2)})}(dx) I_{[a_t=1]} \quad (5)$$

Our next result is concerned with the Markov property. Note that if a jump process is Markovian, it is strong Markovian automatically because of its sample function property.

Theorem 4. The following statements are equivalent:

1°  $X$  is Markovian;

2°  $(T_n, X_{T_n})_{n \geq 0}$  is a homogeneous Markovian chain with state space  $\overline{\mathbb{R}}_+ \times \mathbb{R}$ , and its transition probability  $Q(s, x; dt, dy)$  satisfies the following conditions:

$$1) Q(s, x; dt, dy) = Q(s, x; ]u, \infty] \times \mathbb{R}) Q(u, x; dt, dy) \quad 0 \leq s \leq u \leq t \quad (6)$$

$$2) Q(s, x; ]0, s] \times \mathbb{R}) = Q(s, x; \mathbb{R}_+ \times \{x\}) = 0, \quad s < \infty \quad (7)$$

$$Q(s, x; \{\infty\}, dy) = Q(s, x; \{\infty\} \times \mathbb{R}) \mathcal{E}_{(x)}(dy) \quad (8)$$

$$3) Q(\infty, x; dt, dy) = \mathcal{E}_{(\infty)}(dt) \mathcal{E}_{(x)}(dy) \quad (9)$$

$$3^\circ v(dt, dx) = Q(t, X_{t-}; X_{t-} + dx) \wedge (X_{t-}, dt) \quad (9)$$

where 1)  $Q(t, x; dy)$  is a transition probability from  $\mathbb{R}_+ \times \mathbb{R}$  to  $\mathbb{R}$  and  $Q(t, x; \{x\}) = 0$ ; 2) (i)  $\wedge(x, dt)$  is a  $\sigma$ -finite transition measure from  $\mathbb{R}$  to  $\mathbb{R}_+$  and  $\wedge(x, \{t\}) \leq 1$ ,

(ii) There exist two sequences of Borel functions  $f_n(x)$  and  $g_n(x)$  such that for every  $x$ ,  $\mathbb{R}_+$  is the union of disjoint intervals  $\bigcup_{n=1}^{\infty} [f_n(x), g_n(x)[$ , and for  $t \in [f_n(x), g_n(x)[$

$$\wedge(x, ]f_n(x), t[) < \infty. \quad (10)$$

This problem was firstly discussed by Jacobsen[6] in a slightly different form and under the hypothesis that the state space is denumerable. Gihman and Skorohod [2] essentially showed that the statements 1° and 2° are equivalent, though their proof utilized rather complicated calculation. In fact, one can use the following formulae of jump processes to simplify the calculation. If  $(W_t)_{t \geq 0}$  is an integrable process, then its optional and predictable projections respectively are:

$${}^oW_t = \sum_{n=1}^{\infty} \frac{E(W_t I_{[T_n > t]} | \mathbb{F}_{T_{n-1}})}{E(I_{[T_n > t]} | \mathbb{F}_{T_{n-1}})} I_{[T_{n-1} \leq t < T_n]}$$

and

$$P W_t = \begin{cases} \sum_{n=1}^{\infty} \frac{E(W_t I_{[T_n \geq t]} | \mathbb{F}_{T_{n-1}})}{E(I_{[T_n \geq t]} | \mathbb{F}_{T_{n-1}})} I_{[T_{n-1} < t \leq T_n]}, & t > 0, \\ W_0 & , t = 0. \end{cases}$$

We observe some particular cases. 1) In order that  $X$  is homogeneous Markovian it is necessary and sufficient that  $Q(t, x; dy)$  are independent of  $t$ , and  $\Lambda(x, dt) = q(x)dt$ , with  $q(x) \geq 0$ . Hence we have

$$v(dt, dx) = Q(X_{t-}; X_{t-} + dx)q(X_{t-})dt .$$

This is well-known for the homogeneous Markovian processes with discrete state space (see Jacod[8]). 2) In order that  $X$  is a process with independent increments it is necessary and sufficient that  $Q(t, x; dy)$  and  $\Lambda(x, dt)$  are independent of  $x$ . Hence we have

$$v(dt, dx) = Q(t; dx)dA_t$$

In addition, if  $X$  is stationary, then

$$v(dt, dx) = \lambda Q(dx)dt , \quad \lambda > 0 .$$

These are the results of [9].

2. Predictable representation property. Note that in our case all local martingales are purely discontinuous, and we can deduce the following lemma from the relevant results in Jacod[8].

Lemma 1. Let  $M$  be a local martingale. Then every local martingale  $L$ , with  $L_0 = 0$ , can be represented as a predictable stochastic integral  $H.M$  if and only if the

following conditions are satisfied:

- 1) For every totally inaccessible stopping time  $T$ ,  $[[T]] \subset [\Delta M \neq 0]$ ;
- 2) For every stopping time  $T$ ,  $\underline{F}_T = \underline{F}_{T-} \vee \alpha\{\Delta M_T I_{[T < \infty]}\}$ ;
- 3) There exist two predictable processes  $(\alpha_t^{(i)})$ ,  $i = 1, 2$ , such that  $\Delta M$  equals to  $\alpha^{(1)}$  or  $\alpha^{(2)}$ .

Lemma 2.  $K = [a = 1]$  is the largest predictable set contained in  $D = [\Delta X \neq 0]$ .

Proof. Let  $B$  be a predictable set contained in  $D$ , and  $S$  a predictable stopping time, with  $[[S]] \subset B$ . Then

$$a_S I_{[S < \infty]} = E[I_D(S) I_{[S < \infty]} | \underline{F}_{S-}] = I_{[S < \infty]}.$$

Hence,  $[[S]] \subset K$ , and  $B \subset K$ .  $K \subset D$  is evident.

Proof of theorem 1. No loss generality we can suppose that  $X$  is locally integrable, i.e. its predictable dual projection  $X^p$  exists. Otherwise, we can replace  $X$  by another jump process  $\tilde{X}$  without change of its jump times and natural filtration as follows.

$$\tilde{X} = X_0 + \sum_{n=1}^{\infty} \tilde{\Delta}_n I_{[[T_n, \infty[}}, \quad \tilde{\Delta}_n = \arctg \Delta_n.$$

Then  $\tilde{X}$  is locally integrable, since  $(\tilde{\Delta}_n)_{n \geq 1}$  is bounded.

1°  $\Rightarrow$  2°. Suppose that every local martingale can be represented as a predictable stochastic integral with respect to a local martingale  $M$ . Then  $X - X^p = H \cdot M$ , where  $H$  is a predictable process. By lemma 1 there exist two predictable processes  $(\tilde{\alpha}_t^{(i)})$ ,  $i = 1, 2$ , such that  $\Delta M$  equals to  $\tilde{\alpha}^{(1)}$  or  $\tilde{\alpha}^{(2)}$ . Put

$$\bar{\alpha}^{(i)} = \Delta X^p + H \tilde{\alpha}^{(i)}, \quad i = 1, 2,$$

and

$$\begin{aligned} \alpha^{(1)} &= \bar{\alpha}^{(1)} I_{[|\bar{\alpha}^{(1)}| \geq |\bar{\alpha}^{(2)}|]} + \bar{\alpha}^{(2)} I_{[|\bar{\alpha}^{(1)}| < |\bar{\alpha}^{(2)}|]}, \\ \alpha^{(2)} &= \bar{\alpha}^{(2)} I_{[|\bar{\alpha}^{(1)}| \geq |\bar{\alpha}^{(2)}|]} + \bar{\alpha}^{(1)} I_{[|\bar{\alpha}^{(1)}| < |\bar{\alpha}^{(2)}|]}. \end{aligned}$$

Then  $\Delta X$  equals to  $\alpha^{(1)}$  or  $\alpha^{(2)}$ , and  $|\alpha^{(2)}| \leq |\alpha^{(1)}|$ . Hence we obtain

$$[|\alpha^{(2)}| > 0] \subset [\Delta X \neq 0].$$

Since  $[|\alpha^{(2)}| > 0]$  is predictable, by lemma 2 we have

$$[|\alpha^{(2)}| > 0] \subset [a = 1].$$

Now it is easy to see that for  $n \geq 1$  on the set  $\{T_n < \infty\}$

$$\Delta_n = \Delta X_{T_n} \in \{\alpha_{T_n}^{(1)}, \alpha_{T_n}^{(2)}\}.$$

But on  $\{ a_{T_n} < 1 \}$ ,  $\alpha_{T_n}^{(2)} = 0$ , it must be  $\Delta_m = \alpha_{T_n}^{(1)}$ . On the other hand, since  $\alpha^{(i)}$ ,  $i = 1, 2$ , are predictable, we have  $\alpha_{T_n}^{(i)} \in \underline{F}_{T_n-}$ . So by (2)  $\alpha_{T_n}^{(i)}$  can be represented as

$$\alpha_{T_n}^{(i)} = f_n^{(i)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n) \quad \text{a.s.} \quad i = 1, 2,$$

where  $f_n^{(i)}$ ,  $i = 1, 2$ , are Borel measurable.

2°  $\Rightarrow$  1°. It suffices to verify that the local martingale  $M = X - X^P$  satisfies the conditions in lemma 1.

1) For every totally inaccessible stopping time  $T$ , we have  $[[T]] \subset D$ . Therefore, on the set  $\{ T < \infty \}$ ,  $\Delta X_T \neq 0$ ,  $\Delta X_T^P = 0$ , because  $X^P$  is predictable. This yields  $\Delta M_T \neq 0$ , i.e.  $[[T]] \subset [\Delta M \neq 0]$ .

2) For every stopping time  $T$ , we have  $\Delta X_T^P I_{[T < \infty]} \in \underline{F}_{T-}$ . So by (1)

$$\begin{aligned} \Delta X_T I_{[T < \infty]} &\in \underline{F}_{T-} \vee \sigma\{ \Delta M_T I_{[T < \infty]} \}, \\ \underline{F}_T &= \underline{F}_{T-} \vee \sigma\{ \Delta X_T I_{[T < \infty]} \} = \underline{F}_{T-} \vee \sigma\{ \Delta M_T I_{[T < \infty]} \}. \end{aligned}$$

3) Put

$$\begin{aligned} \tilde{\alpha}^{(1)} &= \sum_{n=1}^{\infty} f_n^{(1)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, t) I_{[[T_{n-1}, T_n]]} \\ \tilde{\alpha}^{(2)} &= I_{[a=1]} \sum_{n=1}^{\infty} f_n^{(2)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, t) I_{[[T_{n-1}, T_n]]} \end{aligned} \quad (11)$$

Then  $\tilde{\alpha}^{(i)}$ ,  $i = 1, 2$ , are predictable and  $\Delta X$  equals to  $\tilde{\alpha}^{(1)}$  or  $\tilde{\alpha}^{(2)}$ . In reality, if  $\Delta X_t = 0$ , it must be  $a_t \leq 1$ , and  $\tilde{\alpha}^{(2)} = 0$ ; if  $\Delta X_t \neq 0$ , there exists an  $n \geq 1$  such that  $t = T_n$ , then  $\Delta X_t = \Delta_n \in \{ f_n^{(i)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n), i = 1, 2 \} = \{ \tilde{\alpha}_{T_n}^{(i)}, i = 1, 2 \} = \{ \tilde{\alpha}_t^{(i)}, i = 1, 2 \}$ . Now set

$$\alpha^{(i)} = -\Delta X^P + \tilde{\alpha}^{(i)}, \quad i = 1, 2,$$

$\alpha^{(i)}$ ,  $i = 1, 2$ , are predictable, and  $\Delta M$  equals to  $\alpha^{(1)}$  or  $\alpha^{(2)}$ .

2°  $\Rightarrow$  3°. For  $n \geq 1$ , put

$$\begin{aligned} P(\Delta_n = f_n^{(i)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n) | \underline{F}_{T_n-}) &= c_n^{(i)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n) \\ c^{(i)} &= \sum_{n=1}^{\infty} c_n^{(i)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, t) I_{[[T_{n-1}, T_n]]}, \quad i = 1, 2. \end{aligned}$$

Then  $c^{(i)}$ ,  $i = 1, 2$ , are predictable, and  $c^{(1)} \geq 0$ ,  $c^{(2)} \geq 0$ ,  $c^{(1)} + c^{(2)} = 1$ . On the set  $\{ T_n < \infty \}$  we have

$$\begin{aligned} P(\Delta_n \in dx | \underline{F}_{T_n-}) &= c_n^{(1)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n) \mathcal{E}(f_n^{(1)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n))(dx) \\ &+ c_n^{(2)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n) \mathcal{E}(f_n^{(2)}(\Delta_{0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n))(dx) I_{[a_{T_n} = 1]}. \end{aligned}$$

By (4) we obtain

$$G(t, dx) = c_t^{(1)} e_{(\alpha_t^{(1)})}(dx) + c_t^{(2)} e_{(\alpha_t^{(2)})}(dx) I_{[a_t=1]}$$

where predictable processes  $\alpha^{(i)}$ ,  $i = 1, 2$ , are defined as above.

3°  $\Rightarrow$  2°. It suffices to see that for every  $n \geq 1$  on the set  $\{T_n < \infty\}$

$$P(\Delta_n \in dx | \mathcal{F}_{T_n^-}) = G(T_n, dx) = c_{T_n}^{(1)} e_{(\alpha_{T_n}^{(1)})}(dx) + c_{T_n}^{(2)} e_{(\alpha_{T_n}^{(2)})}(dx) I_{[a_{T_n}=1]}$$

and to represent  $\alpha_{T_n}^{(i)}$  as  $f_n^{(i)}(\Delta_0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n)$ ,  $i = 1, 2$ .

Corollary 1 ([1]). If for all  $n \geq 1$ ,  $\Delta_n \neq 0 \Rightarrow \Delta_n = 1$ , i.e.  $X$  is a simple point process, then  $X$  has the predictable representation property.

Corollary 2 ([4]). If  $\underline{F}$  is quasi-left-continuous, then  $X$  has the predictable representation property if and only if for every  $n \geq 1$ ,  $\Delta_n = f_n(\Delta_0, T_1, T_2, \dots, T_n)$  a.s., where  $f_n$  is Borel measurable.

Proof. Because of the quasi-left-continuity of  $\underline{F}$ , for every  $n \geq 1$ , on the set  $\{a_{T_n} > 0, T_n < \infty\}$  we have  $\Delta_n = h_n(\Delta_0, T_1, \Delta_1, \dots, T_{n-1}, \Delta_{n-1}, T_n)$  a.s., where  $h_n$  is Borel measurable (see [3] or [5]). Now the corollary can be deduced directly from the statement 2° in theorem 1.

Theorem 2. Let  $(S_n)_{n \geq 1}$  be a sequence of predictable stopping times such that  $D \subset \bigcup_{n=1}^{\infty} \llbracket S_n \rrbracket$  and the graphs  $(\llbracket S_n \rrbracket)_{n \geq 1}$  are disjoint, i.e.  $X$  is accessible. Then  $X$  has the predictable representation property if and only if for every  $n \geq 1$  there exist two  $\mathcal{F}_{S_n^-}$ -measurable variables  $\xi_n^{(i)}$ ,  $i = 1, 2$ , such that on the set  $\{S_n < \infty\}$   $\Delta_{X_{S_n}}$  equals to  $\xi_n^{(1)}$  or  $\xi_n^{(2)}$ . In other words, on the set  $\{S_n < \infty\}$  the conditional distribution of  $\Delta_{X_{S_n}}$  with respect to  $\mathcal{F}_{S_n^-}$  is a two-valued discrete distribution.

The proof of theorem 2 is completely similar to that of theorem 1. It suffices to construct two predictable processes  $\tilde{\alpha}^{(i)}$ ,  $i = 1, 2$ , as follows.

$$\tilde{\alpha}^{(1)} = \sum_{n=1}^{\infty} \xi_n^{(1)} I_{\llbracket S_n \rrbracket}, \quad \tilde{\alpha}^{(2)} = \sum_{n=1}^{\infty} \xi_n^{(2)} I_{\llbracket S_n \rrbracket}$$

instead of (11). In reality, for each  $t$  and  $\omega$ , either  $t = S_n$  for some  $n \geq 1$ ,

$$\Delta_{X_t} = \Delta_{X_{S_n}} \in \{ \xi_n^{(1)}, \xi_n^{(2)} \} = \{ \tilde{\alpha}_{S_n}^{(1)}, \tilde{\alpha}_{S_n}^{(2)} \} = \{ \tilde{\alpha}_t^{(1)}, \tilde{\alpha}_t^{(2)} \},$$

or  $t \in \bigcup_{n=1}^{\infty} \llbracket S_n \rrbracket^c$ ,  $\Delta_{X_t} = 0 = \tilde{\alpha}_t^{(2)}$ . Hence, we still have

$$\Delta_{X_t} \in \{ \tilde{\alpha}_t^{(1)}, \tilde{\alpha}_t^{(2)} \}.$$



Corollary. Let  $X = (X_n)_{n \geq 0}$  be an arbitrary sequence of random variables. Then  $X$  has the predictable representation property if and only if for every  $n \geq 1$ , there exist two  $(X_0, \dots, X_{n-1})$ -measurable variables  $\xi_n^{(i)}$ ,  $i = 1, 2$ , such that  $X_n = \xi_n^{(1)}$  or  $\xi_n^{(2)}$ . In other words, the conditional distribution of  $X_n$  with respect to  $(X_0, \dots, X_{n-1})$  is a two-valued discrete distribution.

In addition, if  $(X_n)_{n \geq 0}$  is an independent sequence, then  $X$  has the predictable representation property if and only if each of  $(X_n)_{n \geq 1}$  has a two-valued discrete distribution.

Proof. Define a jump process

$$X_t = X_0 + \sum_{n=1}^{\infty} (X_n - X_{n-1}) I_{[n \leq t]}$$

and take  $S_n = n$ . The conclusions follow immediately from theorem 2.

Though the corollary of theorem 2 is rather banal, it motivated the following general result on the processes with independent increments (not necessarily stochastically continuous) (see [4]).

Theorem 3. Suppose that  $X = (X_t)_{t \geq 0}$  is a process with independent increments, and with r.c.l.l. trajectories. Let  $(\alpha, \beta, \nu)$  be the local characteristics of  $X$ . Then  $X$  has the predictable representation property if and only if

$$1) \nu(dt, dx) = \{c_t^{(1)} e_{f_t^{(1)}}(dx) + c_t^{(2)} e_{f_t^{(2)}}(dx) I_{[\nu(\{t\} \times \mathbb{R}) > 0]}\} d\Lambda_t,$$

where  $c^{(i)}$ ,  $f^{(i)}$ ,  $i = 1, 2$ , are Borel measurable functions, with  $c^{(1)} \geq 0$ ,  $c^{(2)} \geq 0$ ,  $c^{(1)} + c^{(2)} = 1$ , and  $d\Lambda_t$  is a  $\sigma$ -finite measure on  $\mathbb{R}_+$ ;

2)  $d\beta_t$  and  $d\Lambda_t$  are mutually singular.

Note that  $[\nu(\{t\} \times \mathbb{R}) > 0]$  is the set of the fixed discontinuous points of  $X$ , only on this set the jumps of  $X$  can take two possible values.

3. Markov property. We turn to Markov property of jump processes and complete the demonstration of theorem 4 by proving that the statements 2° and 3° are equivalent.

2°  $\Rightarrow$  3°. For  $s \leq t$ , put

$$q(s, x, t) = Q(s, x; ]t, \infty] \times \mathbb{R}) \quad .$$

$q(s, x, \cdot)$  is right-continuous and monotonely decreasing, and by (6) it satisfies the following functional equation:

$$\begin{aligned} q(s, x, t) &= q(s, x, u)q(u, x, t) & s \leq u \leq t \\ q(s, x, s) &= 1 \end{aligned} \quad (12)$$

Denote  $\tau_g(x) = \inf \{ t > s : q(s, x, t) = 0 \}$ . From (12) it is facile to get

$$\begin{aligned} 1) \tau_g(x) &> s; \\ 2) q(s, x, u) &> 0, u \in [s, \tau_g(x)[; \\ 3) q(s, x, u) &= 0, u \in [\tau_g(x), \infty[. \end{aligned} \quad (13)$$

We can decompose  $R_+$  into a series of disjoint intervals:  $R_+ = \bigcup_{n=1}^{\infty} [f_n(x), g_n(x)[$  such that for arbitrary two points  $s$  and  $t$  ( $s < t$ ),  $q(s, x, t) > 0$  if  $s$  and  $t$  belong to the same interval, and  $q(s, x, t) = 0$  if  $s$  and  $t$  belong to different intervals.

In fact, for  $x$  fixed we may classify the points of  $R_+$  as follows. For  $s < t$ , we stipulate that  $s$  and  $t$  belong to the same class  $C_\alpha(x)$  if and only if  $q(s, x, t) > 0$ . Because of (12) there is no ambiguity. It suffices to prove that each class  $C_\alpha(x)$  is an interval  $[f_\alpha(x), g_\alpha(x)[$ , since the number of disjoint intervals on  $R_+$  is at most denumerable. From (13) the proof is straightforward. We observe that if  $s$  and  $t$  belong to  $C_\alpha(x)$  and  $s < t$ , then  $[s, t] \subset C_\alpha(x)$ . Set  $f_\alpha(x) = \inf C_\alpha(x)$ ,  $g_\alpha(x) = \sup C_\alpha(x)$ , we get

$$[f_\alpha(x), g_\alpha(x)[ \subset C_\alpha(x) \subset [f_\alpha(x), g_\alpha(x)].$$

It remains to show  $f_\alpha(x) \in C_\alpha(x)$  and  $g_\alpha(x) \notin C_\alpha(x)$  if  $g_\alpha(x) < \infty$ . Take  $u \in [f_\alpha(x), g_\alpha(x)[$  such that  $q(f_\alpha(x), x, u) > 0$ . Then by (12) for every  $t \in C_\alpha(x)$ ,  $q(f_\alpha(x), x, t) > 0$ , and this yields  $f_\alpha(x) \in C_\alpha(x)$ . Now suppose  $g_\alpha(x) < \infty$ . there exists  $u > g_\alpha(x)$  such that  $q(g_\alpha(x), x, u) > 0$ . If  $g_\alpha(x) \in C_\alpha(x)$ , then  $u \in C_\alpha(x)$ . This contradicts to the fact that  $g_\alpha(x)$  is the supremum of  $C_\alpha(x)$ .

Furthermore, we can consider  $f_n(x)$  and  $g_n(x)$  to be measurable. In fact, we need only to arrange those intervals, whose lengths are more than  $\frac{1}{n}$  and not more than  $\frac{1}{n-1}$ , and the number of such intervals in every finite time interval is finite. Set

$$\begin{aligned} a_0^{(n)}(x) &= b_0^{(n)}(x) = 0, \\ a_m^{(n)}(x) &= \inf \left\{ t > b_{m-1}^{(n)}(x) : q(t, x, t + \frac{1}{n}) > 0, q(t, x, t + \frac{1}{n-1}) = 0 \right\}, \\ b_m^{(n)}(x) &= \inf \left\{ t > a_m^{(n)}(x) : q(a_m^{(n)}(x), x, t) = 0 \right\}. \end{aligned} \quad (14)$$

Then  $R_+ = \bigcup_{n=1}^{\infty} \bigcup_{m=1}^{\infty} [a_m^{(n)}(x), b_m^{(n)}(x)[$ . Because  $q(t, x, t + \delta)$  ( $\delta > 0$ ) and  $q(a_m^{(n)}(x), x, t)$  are right-continuous in  $t$ , the infimums in (14) can be taken over the

rational numbers. Hence,  $a_m^{(n)}(x)$  and  $b_m^{(n)}(x)$  are measurable. Taking away empty intervals and rearrange properly, we obtain the decomposition  $R_+ = \bigcup_{n=1}^{\infty} [f_n(x), g_n(x)[$  with measurable end point functions.

Put

$$\Lambda_n(x, dt) = \frac{q(f_n(x), x; dt)}{q(f_n(x), x; [t, \infty])}, \quad q(s, x; dt) = Q(s, x; dt, R)$$

$$\Lambda(x, dt) = \sum_{n=1}^{\infty} \Lambda_n(x, dt).$$

Note that the support of  $\Lambda_n(x, dt)$  is  $[f_n(x), g_n(x)]$  and  $\Lambda_n(x, \{t\}) \leq 1$ ,

$$\Lambda_n(x, [f_n(x), u]) < \infty, \quad u \in [f_n(x), g_n(x)[.$$

So  $\Lambda(x, dt)$  is well defined and satisfies the conditions demanded in the statement 2:

Take

$$Q_n(t, x; dy) = \frac{Q(f_n(x), x; dt, dy)}{q(f_n(x), x; dt)}$$

as the Radon-Nikodym derivative of  $Q(f_n(x), x; dt, dy)$  with respect to  $q(f_n(x), x, dt)$  such that it is a transition probability and vanishes for  $t \in [f_n(x), g_n(x)]$ . Similarly we define

$$Q(t, x; dy) = \sum_{n=1}^{\infty} Q_n(t, x; dy),$$

which is a transition probability from  $R_+ \times R$  to  $R$ .

Now we verify the formula (7). Fix  $n \geq 1$ . On the set  $\{T_{n-1} \in [f_k(X_{T_{n-1}}), g_k(X_{T_{n-1}})]\}$  we have  $q(T_{n-1}, X_{T_{n-1}}, [g_k(X_{T_{n-1}}), \infty]) = 0$ , so  $T_n \leq g_k(X_{T_{n-1}})$ , i.e.

$$]T_{n-1}, T_n] \subset [f_k(X_{T_{n-1}}), g_k(X_{T_{n-1}})].$$

On the other hand, by (10) for any  $u \in [f_n(x), g_n(x)[$  we have

$$\frac{q(u, x; dt)}{q(u, x; [t, \infty])} = \Lambda_n(x, dt), \quad t \geq u,$$

particularly,

$$\frac{q(T_{n-1}, X_{T_{n-1}}; dt)}{q(T_{n-1}, X_{T_{n-1}}; [t, \infty])} = \Lambda_k(X_{T_{n-1}}, dt).$$

Hence,

$$\frac{Q(T_{n-1}, X_{T_{n-1}}; dt, X_{T_{n-1}} + dx)}{q(T_{n-1}, X_{T_{n-1}}; [t, \infty])} I_{[T_{n-1} < t \leq T_n]}$$

$$= Q_k(t, X_{T_{n-1}}; X_{T_{n-1}} + dx) \Lambda_k(X_{T_{n-1}}, dt) I_{[T_{n-1} < t \leq T_n]}$$

$$= Q(t, X_{t-}; X_{t-} + dx) \Lambda(X_{t-}, dt) I_{[\tau_{n-1} < t \leq \tau_n]} .$$

According to (3) and utilizing the Markov property of  $(T_n, X_{T_n})_{n \geq 0}$  we get

$$\begin{aligned} \nu(dt, dx) &= \sum_{n=1}^{\infty} \frac{Q(\tau_{n-1}, X_{\tau_{n-1}}; dt, X_{\tau_{n-1}} + dx)}{q(\tau_{n-1}, X_{\tau_{n-1}}; [t, \infty])} I_{[\tau_{n-1} < t \leq \tau_n]} \\ &= Q(t, X_{t-}; X_{t-} + dx) \Lambda(X_{t-}, dt) . \end{aligned}$$

Remark. If  $(X_t)_{t \geq 0}$  is a homogeneous Markovian process, the functions  $q(s, x, t)$  are only dependent of  $t - s$ :  $q(s, x, t) = q(t - s, x)$ , and equation (12) becomes

$$q(s + t, x) = q(s, x)q(t, x) , \quad s, t \geq 0 .$$

Immediately, we have  $q(t, x) = e^{-q(x)t}$ , hence  $\Lambda(x, dt) = q(x)dt$ , and  $Q(t, x; dy)$  is independent of  $t$ . At the same time, since  $\Lambda(x, dt)$  is continuous in  $t$ ,  $X$  is quasi-left-continuous, i.e. all  $(T_n)_{n \geq 1}$  are totally inaccessible.

$3^\circ \Rightarrow 2^\circ$ . According to Doleans-Dade's exponential formula, we define

$$\begin{aligned} q(s, x, t) &= e^{-\Lambda^c(x, ]s, t[ \wedge g_n(x))} \prod_{s < u \leq t} \Lambda_{g_n(x)}(1 - \Lambda(x, \{u\})), & s \leq t , \quad (15) \\ Q(s, x; dt, dy) &= Q(u, x; dy)q(s, x; du) I_{]s, g_n(x)[}(u), & s \in [f_n(x), g_n(x)[ \end{aligned}$$

where  $\Lambda^c(x, dt)$  is the continuous part of  $\Lambda(x, dt)$ . It is facile to verify that  $Q(s, x; dt, dy)$  defined in (15) together with (7), (8) constitutes a transition probability and satisfies the condition (6).

Now we can construct a jump process  $\bar{X}$  such that the corresponding chain  $(\bar{T}_n, \bar{X}_{\bar{T}_n})_{n \geq 0}$  is homogeneous Markovian with  $Q(s, x; dt, dy)$  as its transition probability, and  $\bar{X}_0$  has the same law as  $X_0$ . Then from the proof  $2^\circ \Rightarrow 3^\circ$ , the corresponding predictable dual projection  $\bar{\nu}$  has the same form as  $\nu$

$$\bar{\nu}(dt, dx) = Q(t, \bar{X}_{t-}; \bar{X}_{t-} + dx) \Lambda(\bar{X}_{t-}, dt) .$$

Therefore,  $\bar{X}$  has the same law as  $X$ . This implies that  $(T_n, X_{T_n})_{n \geq 0}$  has the same law as  $(\bar{T}_n, \bar{X}_{\bar{T}_n})_{n \geq 0}$ . Hence,  $(T_n, X_{T_n})_{n \geq 0}$  is a homogeneous Markovian chain.

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